DIRAC-TYPE PROBLEM OF RAINBOW MATCHINGS AND HAMILTON CYCLES IN RANDOM GRAPHS

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ABSTRACT. Given a family of graphs G_1, \ldots, G_n on the same vertex set [n], a rainbow Hamilton cycle is a Hamilton cycle on [n] such that each G_i contributes exactly one edge. We prove that if G_1, \ldots, G_n are independent samples of G(n, p) on the same vertex set [n], then for each $\varepsilon > 0$, whp, every collection of spanning subgraphs $H_i \subseteq G_i$, with $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)np$, admits a rainbow Hamilton cycle. A similar result is proved for rainbow perfect matchings in a family of n/2 graphs on the same vertex set [n].

Our method is likely to be applicable to further problems in the rainbow setting, in particular, we illustrate how it works for finding a rainbow perfect matching in the k-partite k-uniform hypergraph setting.

1. INTRODUCTION

2 1.1. Dirac-type problems. Arguably the two most studied objects in graph theory are *perfect* 3 matchings and Hamilton cycles. A perfect matching in a graph G = (V, E) is a collection of vertex-4 disjoint edges which covers V, and a Hamilton cycle is a cycle passing through all the vertices of 5 G. As opposed to the problem of finding a perfect matching (if one exists) in a graph G which has 6 efficient (polynomial-time) resolutions, the analogous problem for Hamilton cycles is listed as one 7 of the NP-hard problems by Karp [14]. Therefore, as one cannot hope to find a Hamilton cycle 8 efficiently, it is natural to study sufficient conditions which guarantee its existence.

One of the first results of this type is the celebrated theorem by Dirac [11], which states that every 9 graph on $n \geq 3$ vertices with minimum degree $\frac{n}{2}$ is Hamiltonian, that is, contains a Hamilton cycle 10 (and in particular, if n is even, then it also contains a perfect matching). While Dirac's theorem is 11 sharp in general, one would like to find sufficient conditions for sparser graphs. A natural candidate 12 to begin with is a typical graph sampled from the binomial random graph model G(n,p). That 13 is, a graph G on vertex set [n], where each (unordered) pair is being sampled as an edge with 14 probability p, independently. In 1960, Erdős and Rényi raised a question of what the threshold 15 probability of Hamiltonicity in random graphs is. This question attracted a lot of attention in 16 the past few decades. After a series of efforts by various researchers, including Korshunov [16] 17 and Pósa [26], the problem was finally solved by Komlós and Szemerédi [15] and independently by 18 Bollobás [4], who proved that if $p \ge (\log n + \log \log n + \omega(1))/n$, where $\omega(1)$ tends to infinity with 19 n arbitrarily slowly, then the probability of the random graph G(n, p) being Hamiltonian tends 20 to 1 (we say such an event happens with high probability, or whp for brevity). This result is best 21 possible since for $p \leq (\log n + \log \log n - \omega(1))/n$ where are vertices of degree at most one in 22 G(n,p) (see, e.g. [5]). An even stronger result was given by Bollobás [4]. He showed that for the 23 random graph process, the hitting time for Hamiltonicity is exactly the same as the hitting time 24 for having minimum degree 2, that is, who the very edge which increases the minimum degree to 25 2 also makes the graph Hamiltonian. 26

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In this paper, we take advantage of the study of the local resilience in random graphs and random digraphs, which was first introduced by Sudakov and Vu [27]. Roughly speaking, the local resilience is the largest proportion of edges that one can delete from every vertex in a given graph G satisfying a property \mathcal{P} , such that the resulting (sub)graph still satisfies \mathcal{P} . We shall use resilience results on perfect matchings in random bipartite graphs due to Sudakov and Vu [27] and Hamiltonicity in random digraphs by Montgomery [22], see Lemma 2.9 and Lemma 2.13, respectively.

1.2. A rainbow setting. In recent years, rainbow structures in graph systems have received a lot of attention [1, 2, 6, 7, 8, 9, 10, 13, 17, 18, 19, 20, 23]. Formally, given a family of (hyper)graphs $\mathcal{G} = \{G_1, \ldots, G_m\}$ defined on the same vertex set, a copy of an *m*-edge (hyper)graph *H* is called *rainbow* if $E(H) \subseteq \bigcup_{i \in [m]} E(G_i)$ and $|E(H) \cap E(G_i)| = 1$ for every $i \in [m]$.

A rainbow version of the Dirac-type problems in systems of graphs was conjectured by Aharoni 37 et al.[1]: let G_1, \ldots, G_n be a system of graphs on the same vertex set V = [n] with minimum 38 degree $\delta(G_i) \geq n/2$ for each $i \in [n]$, there exists a rainbow Hamilton cycle. Cheng, Wang and Zhao 39 [10] verified the conjecture asymptotically and Joos and Kim [13] proved the full conjecture. Very 40 recently, Bradshaw, Halasz, and Stacho [7] strengthened the result by showing that the system of 41 graphs actually admit exponentially many rainbow Hamilton cycles under the same assumptions. 42 Bradshaw [6] generalized the Dirac-type result for Hamiltonicity of bipartite graphs by Moon and 43 Moser [24] to the rainbow setting. 44

Another interesting structure to consider is the perfect matching. In the rainbow setting we are given n/2 graphs $G_1, \ldots, G_{n/2}$ on the same vertex set [n], and we are seeking for a rainbow perfect matching, namely, a perfect matching that consists of exactly one edge from each G_i .

In this note, we give Dirac-type results for rainbow perfect matchings and rainbow Hamilton cycles in random graphs. Our main results read as follows.

Theorem 1.1. Let $\varepsilon > 0$ and $p = \omega\left(\frac{\log n}{n}\right)$ where *n* is even. Suppose $G_1, \ldots, G_{n/2}$ are independent samples of G(n,p) on the same vertex set V = [n]. Then, whp we have that for every spanning subgraphs $H_i \subseteq G_i$, $1 \le i \le n/2$, with $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)np$, the family $\{H_1, \ldots, H_{n/2}\}$ admits a rainbow perfect matching.

Theorem 1.2. Let $\varepsilon > 0$ and $p = \omega\left(\frac{\log n}{n}\right)$. Suppose G_1, \ldots, G_n are independent samples of 55 G(n,p) on the same vertex set V = [n]. Then, whp we have that for every spanning subgraphs 56 $H_i \subseteq G_i, 1 \le i \le n$, with $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)np$, the family $\{H_1, \ldots, H_n\}$ admits a rainbow Hamilton 57 cycle.

1.3. Notation. Given a graph G and $X \subseteq V(G)$, let $N(X) = \bigcup_{x \in X} N(x)$. For two subsets $X, Y \subseteq V(G)$ we define $E_G(X, Y)$ to be the set of all edges $xy \in E(G)$ with $x \in X$ and $y \in Y$, and set $e_G(X, Y) := |E(X, Y)|$ (the subscript G will be omitted whenever there is no risk of confusion). Moreover, G[X, Y] is defined by a graph with vertex set $X \cup Y$ and edge set $E_G(X, Y)$. When $x \in V(G), d_G(x)$ is the *degree* of x in G. For a graph $H, X \subset V(H)$, and a vertex $v \in V(H)$, we define $d_H(v, X) = |\{uv \in E(H) : u \in X\}|$. In particular, if $X = \{u\}$ for some vertex u, then we write $d(v, u) := d(v, \{u\})$.

For a graph G, we denote by $\delta(G)$ as its minimum degree. For a digraph D, we denote $\delta^+(D), \delta^-(D)$ as its minimum out-degree and minimum in-degree, respectively. Moreover, let

$$\delta^0(D) = \min\{\delta^+(D), \delta^-(D)\}.$$

If $f(n)/g(n) \to 0$ as $n \to \infty$, then we say $g(n) = \omega(f(n))$ and f(n) = o(g(n)). If there exists a constant C for which $f(n) \leq Cg(n)$ for all n, then we say f(n) = O(g(n)) and $g(n) = \Omega(f(n))$. If f = O(g(n)) and $f(n) = \Omega(g(n))$, then we say that $f(n) = \Theta(g(n))$. The random graph G(n, p)has vertex set $[n] = \{1, \ldots, n\}$ and edges chosen independently at random with probability p.

The random digraph D(n, p) has vertex set $[n] = \{1, \ldots, n\}$ and directed edges (u, v), which is an 71 ordered pair of vertices, chosen independently at random with probability p. 72

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2. Preliminary Results

In this section we first present some useful tools for the proofs, and then introduce our key 74 auxiliary graphs and how to use them. 75

2.1. Chernoff's inequalities. We will use the following well-known bound on the upper and lower 76 tails of the binomial distribution, which is given by Chernoff (see Appendix A in [3]). 77

Lemma 2.1 (Chernoff's inequality). Let $X \sim Bin(n, p)$ and let $\mathbb{E}[X] = \mu$. Then 78

- 79
- Pr[X < (1 − a)µ] < e^{-a²µ/2} for every a > 0;
 Pr[X > (1 + a)µ] < e^{-a²µ/3} for every 0 < a < 3/2. 80

Remark 2.2. Chernoff's inequalities also hold when X is hypergeometrically distributed with mean 81 μ. 82

The following simple bound is also useful in our proof. 83

Lemma 2.3. Let $X \sim Bin(m,q)$. Then, for all k we have

$$\Pr[X \ge k] \le \left(\frac{emq}{k}\right)^k.$$

Proof. Indeed, note that

$$\Pr[X \ge k] \le \binom{m}{k} q^k \le \left(\frac{emq}{k}\right)^k$$

as desired. 84

2.2. Talagrand-type inequality. Our main probabilistic tool is the following concentration in-85 equality of McDiarmid [21]. 86

- **Theorem 2.4.** Given a set S of size m, we let Sym(S) denote the set of all m! permutations of S. 87
- Let $\{B_1,\ldots,B_k\}$ be a family of finite non-empty sets, and let $\Omega = \prod_i \operatorname{Sym}(B_i)$. Let $\pi = \{\pi_1,\ldots,\pi_k\}$ 88

be a family of independent permutations, such that for $i, \pi_i \in \text{Sym}(B_i)$ is chosen uniformly at 89 random. 90

Let c and r be constants, and suppose that a nonnegative real-valued function h on Ω satisfies 91 the following conditions for each $\pi \in \Omega$. 92

- (1) Swapping any two elements in any π_i can change the value of h by at most 2c. 93
- (2) If $h(\pi) = s$, there exists a set $\pi_{proof} \subseteq \pi$ of size at most rs, such that $h(\pi') \geq s$ for any 94 $\pi' \in \Omega$ where $\pi' \supseteq \pi_{proof}$. 95

Then for each $t \ge 0$ we have

$$\Pr[h \le M(h(\pi)) - t] \le 2 \exp\left(-\frac{t^2}{16rc^2M}\right)$$

2.3. Typical properties of graphs. In this section, we collect some useful properties of a typical 96 sequence of independent samples of G(n, p), which are regarded as "colors", on the same vertex set 97 V = [n]. First, we show that the degrees are concentrated. 98

Lemma 2.5. Let $\varepsilon > 0$ and let $N \leq n^2$. Let G_1, \ldots, G_N be independent samples of G(n, p) on the same vertex set V = [n]. Then, why we have

$$(1 - \varepsilon)np \le \delta(G_c) \le \Delta(G_c) \le (1 + \varepsilon)np$$

holds for all $c \in [N]$ provided that $p = \omega(\frac{\log n}{n})$. 99

Proof. Fix some vertex $u \in [n]$ and some color $c \in [N]$. Observe that $d_{G_c}(u) \sim Bin(n-1,p)$, and therefore $\mu := \mathbb{E}[d_{G_c}(u)] = (n-1)p$. Hence, since $p = \omega(\frac{\log n}{n})$, by Lemma 2.1 we obtain that

$$\Pr[d_{G_c}(u) \notin (1 \pm \varepsilon)\mu] \le 2 \exp\left(-\frac{\varepsilon^2 \mu}{3}\right) = o\left(\frac{1}{n^3}\right).$$

Taking a union bound over all vertices $u \in [n]$ and all colors $c \in [N]$, we conclude that

$$\Pr[\exists u \in [n], \exists c \in [N] \text{ s.t. } d_{G_c}(u) \notin (1 \pm \varepsilon)\mu] = o(1).$$

100 This completes the proof.

Next, we show that given n/2 graphs $H_1, \ldots, H_{n/2}$, if we take a random equipartition of [n], then whp the corresponding bipartite subgraphs of H_i have the "correct" degrees.

Lemma 2.6. For every $\varepsilon > 0$ there exists $C := C(\varepsilon)$ for which the following holds. Let m = n/2. Let H_1, \ldots, H_m be graphs on the same vertex set V = [n], where n is a sufficiently large even integer. Suppose that $\delta(H_c) \ge C \log n$ for all $c \in [m]$. Then, a (1 - o(1))-fraction of the partitions $V = V_1 \cup V_2$ into sets of size m satisfy the following property: For every vertex $u \in V$ and $c \in [m]$, and for i = 1, 2 we have

$$d_{H_c}(u, V_i) \in (1 \pm \varepsilon) \cdot \frac{d_{H_c}(u)}{2}.$$

Proof. Consider a random partition $V = V_1 \cup V_2$ into sets both of size m. For some fixed vertex $u \in [n]$ and some fixed $c \in [m]$, note that $d_{H_c}(u)$ is hypergeometrically distributed with expected value $\frac{d_{H_c}(u)}{2}$. Therefore, by Lemma 2.1 we obtain that

$$\Pr\left[d_{H_c}(u, V_i) \notin (1 \pm \varepsilon) \cdot \frac{d_{H_c}(u)}{2}\right] \le 2e^{-\varepsilon^2 \frac{d_{H_c}(u)}{2}/3} \le 2e^{-3\log n} = 2n^{-3},$$

where the last inequality holds for a large enough C.

By applying a union bound over all possible u's and i's, we obtain that the probability of having such a vertex u and an index i is at most $(n^2/2) \cdot 2n^{-3} = n^{-1}$. This completes the proof.

2.4. Auxiliary graphs and proof ideas. In this section we define some auxiliary graphs/digraphs
 that are going to play key roles in the proofs of our main theorems.

108 2.4.1. Perfect matchings. Our proof uses an auxiliary bipartite graph as follows.

Definition 2.7. Let n be an even integer. Let $H'_1, \ldots, H'_{n/2}$ be bipartite graphs on the same vertex set V = [n], each of which has the same bipartition $V = V_1 \cup V_2$ with $|V_1| = |V_2| = n/2$. By relabeling the vertices (if necessary), we may assume $V_1 = [n/2]$. Given a permutation $\pi : V_1 \to V_1$, the auxiliary bipartite graph $B_{\pi} := B_{\pi}(H'_1, \ldots, H'_{n/2})$ is constructed as follows: the parts of B_{π} are V_1 and V_2 ; the edge set consists of all pairs $(i, j) \in V_1 \times V_2$ such that $ij \in E(H'_{\pi(i)})$.

Remark 2.8. Observe that a perfect matching in B_{π} corresponds to a rainbow perfect matching in the family $H'_1, \ldots, H'_{n/2}$. Indeed, every edge $\{i, j\}$ in B_{π} with $i \in V_1$ and $j \in V_2$ corresponds to an edge $\{i, j\}$ in $H'_{\pi(i)}$, and since π is a permutation of V_1 , a perfect matching of B_{π} uses exactly one edge from each H'_i .

We also need the following result of Sudakov and Vu [27] on local resilience of perfect matchings in random bipartite graphs, whose proof can be found in the proof of [27, Theorem 3.1]. Let V_1 and V_2 be disjoint vertex sets each of size n/2, where $n \in 2\mathbb{N}$. A random bipartite graph B(n, p)defined on the partition $V_1 \cup V_2$ is a bipartite graph such that given $(i, j) \in V_1 \times V_2$, $ij \in E(B(n, p))$ with probability p and all pairs ij are chosen independent of each other.

123 Lemma 2.9 (Sudakov and Vu [27]). Let $\varepsilon > 0$ and $p = \omega\left(\frac{\log n}{n}\right)$ where n is even. Let V_1 and V_2 124 be disjoint vertex sets each of size n/2. Then, whp a spanning subgraph $H \subseteq B(n,p)$ defined on 125 $V_1 \cup V_2$ such that $\delta(H) \ge (\frac{1}{2} + \varepsilon)\frac{np}{2}$ contains a perfect matching.

We first demonstrate how to use the auxiliary bipartite graph to prove the following bipartition version of Theorem 1.1.

Theorem 2.10. Let $\varepsilon > 0$ and $p = \omega\left(\frac{\log n}{n}\right)$ where *n* is even. Suppose $G_1, \ldots, G_{n/2}$ are independent samples of B(n,p) defined on $V_1 \cup V_2$, each of size n/2. Then, whp we have that for every spanning (bipartite) subgraphs $H_i \subseteq G_i$, $1 \le i \le n/2$, with $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)\frac{np}{2}$, the family $\{H_1, \ldots, H_{n/2}\}$ admits a rainbow perfect matching.

Here is an outline of our proof. Given a set V of $n \in 2\mathbb{N}$ vertices and a balanced bipartition $V = V_1 \cup V_2$, fix a permutation $\pi : V_1 \to V_1$. If we expose n/2 independent samples of B(n, p) on $V_1 \cup V_2$, denoted by $G_1, \ldots, G_{n/2}$, then we have that the bipartite graph $G_{\pi} = B_{\pi}(G_1, \ldots, G_{n/2})$ defined in Definition 2.7 is also a random bipartite graph.

Then, given $(i, j) \in V_1 \times V_2$, $ij \in G_{\pi}$ if and only if $ij \in E(G_{\pi(i)})$, which happens with probability 136 precisely p and is independent with all other pairs in $V_1 \times V_2$. In particular, where G_{π} is resilient 137 for perfect matching – by Lemma 2.9. That is, whp every spanning subgraph H of G_{π} with 138 $\delta(H) \geq (\frac{1}{2} + \frac{\varepsilon}{2})\frac{np}{2}$ contains a perfect matching. However, once the independent samples are exposed, 139 there might be some "bad" permutations π such that there exists $i \in V_1$ satisfying $ij \notin E(G_{\pi(i)})$ for 140 most $j \in V_2$. Fortunately, we can prove that almost every π is not bad by Markov's inequality. On 141 the other hand, we shall show that (Lemma 2.16) if π is a uniformly random permutation of V_1 and 142 $H_1, \ldots, H_{n/2}$ are graphs such that $H_i \subseteq G_i$ and $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)\frac{np}{2}$, then whp the bipartite graph 143 $B_{\pi} = B_{\pi}(H_1, \ldots, H_{n/2})$ satisfies that $B_{\pi} \subseteq G_{\pi}$ and $\delta(B_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})\frac{np}{2}$. Therefore, by the resilience 144 property of G_{π} we conclude that H_{π} has a perfect matching, which gives rise to a rainbow perfect 145 matching of the family $H_1, \ldots, H_{n/2}$. This will prove Theorem 2.10. 146

To derive Theorem 1.1, it suffices to show that there is a balanced bipartition of the vertex set such that the bipartite subgraphs of our graphs inherits the degree condition, which is proved in Lemma 2.6.

2.4.2. Directed Hamilton cycles. For Hamiltonicity we will need to construct an auxiliary digraph
 as follows.

Definition 2.11. Let H'_1, \ldots, H'_n be graphs on the same vertex set V = [n], Given a permutation 153 $\pi: V \to V$, the auxiliary digraph $D_{\pi} := D_{\pi}(H'_1, \ldots, H'_n)$ is constructed as follows:

154 V is the vertex set of D_{π} and for any vertex $i, j \in V, (i, j) \in E(D_{\pi})$ if and only if $ij \in E(H'_{\pi(i)})$.

Remark 2.12. Observe that a directed Hamilton cycle in the auxiliary digraph D_{π} corresponds to a rainbow Hamilton cycle in the family H'_1, \ldots, H'_n . Indeed, a directed Hamilton cycle in D_{π} is a directed Hamilton cycle in K_n whose edges (i, j) belongs to distinct $H'_{\pi(i)}$ since π is a permutation.

We also need the following result on local resilience of Hamiltonicity in random digraphs due to Montgomery [22] (in fact, Montgomery proved a way stronger result but the following is enough for our needs).

Lemma 2.13. Let $\varepsilon > 0$. Then whp a spanning subdigraph $D \sim D(n, p)$ defined on V = [n] such that $\delta^0(D) \ge (\frac{1}{2} + \varepsilon)np$ contains a Hamilton cycle, provided that $p = \omega(\frac{\log n}{n})$.

Similar to what we did in the previous subsection, we explain how this auxiliary digraph works. Given a set V of n vertices, fix a permutation $\pi: V \to V$. If we expose n independent samples of G(n,p) on V, denoted by G_1, \ldots, G_n , then the digraph $G_{\pi} = D_{\pi}(G_1, \ldots, G_n)$ defined in Definition 2.11 is a random digraph. Indeed, given $(i, j) \in V^2$, then $ij \in G_{\pi}$ if and only if $(i, j) \in E(G_{\pi(i)})$, which happens with probability precisely p and is independent with all other pairs in V^2 . In particular, G_{π} is resilient for Hamiltonicity – by Lemma 2.13, whp any spanning subdigraph H of G_{π} with $\delta^0(H) \ge (\frac{1}{2} + \frac{\varepsilon}{2})np$ contains a directed Hamilton cycle. On the other hand, we shall show that (Lemma 2.17) if π is a uniformly random permutation of V and H_1, \ldots, H_n are graphs used that $H_i \subseteq G_i$ and $\delta(H_i) \ge (\frac{1}{2} + \varepsilon)np$, then whp the digraph $D_{\pi} = D_{\pi}(H_1, \ldots, H_n)$ satisfies that $D_{\pi} \subseteq G_{\pi}$ and $\delta(D_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})np$. Therefore, by the resilience property of G_{π} , we conclude that D_{π} has a directed Hamilton cycle, which gives rise to a rainbow Hamilton cycle of the family H_1, \ldots, H_n . This will prove Theorem 1.2.

175 2.5. Most B_{π} 's have large minimum degree. In this section we prove that given a balanced 176 partition $[n] = V_1 \cup V_2$ and bipartite graphs H'_1, \ldots, H'_m with a common bipartition $V_1 \cup V_2$ (with 177 m = n/2) and large minimum degrees, the resulting auxiliary graph B_{π} also has large minimum 178 degree whp where π is a uniformly random permutation. The proof is a special case of Lemma 13 179 in [12], but we include it for completeness.

Lemma 2.14. Let $0 < \alpha < \frac{1}{2}$ and let $n \in 2\mathbb{N}$ be sufficiently large. Let m = n/2. Let H'_1, \ldots, H'_m be bipartite graphs on the same vertex set V = [n] with the same parts $V = V_1 \cup V_2$ of the same size m, where $V_1 = [m]$. Suppose that $\delta^*(H'_c) \geq \frac{200}{\alpha^2}$ for all $c \in [m]$. Let π be a uniformly random permutation on V_1 and $\mu_i = \mathbb{E}[d_{B_{\pi}}(i)]$. Then, for every $j \in V_2$, we have

$$M_j := M(d_{B_\pi}(j)) \in (1 \pm \alpha)\mu_j.$$

Remark 2.15. The above lemma allows us to use μ_j instead of M_j in Theorem 2.4 when it is applied to $d_{B_{\pi}}(j)$.

182 Proof. Consider B_{π} , where π is a uniformly random permutation on V_1 . Let j be some vertex in 183 V_2 . Let $\mu_j = \mathbb{E}[d_{B_{\pi}}(j)]$ and $\sigma^2 = \operatorname{Var}(d_{B_{\pi}}(j))$. Moreover, for each $i \in V_1$, we define an indicator 184 random variable $\mathbb{1}_i$, where $\mathbb{1}_i = 1$ if $\{i, j\} \in E(H'_{\pi(i)})$. Observe that $d_{B_{\pi}}(j) = \sum_{i \in V_1} \mathbb{1}_i$.

Applying Chebyshev's inequality, we have

$$\Pr[|d_{B_{\pi}}(j) - \mu_j| \ge \alpha \mu_j] \le \frac{\sigma^2}{\alpha^2 \mu_j^2}$$

If we can show that $\sigma^2 \leq \frac{\alpha^2 \mu_j^2}{100}$, then the result follows. Indeed, with probability at least 99/100 we have that $d_{B_{\pi}}(j) \in (1 \pm \alpha) \mu_j$ and thus we conclude that the median M_j also lies in this interval. Now the remaining part is to prove the desired inequality by computing $\mu_j = \mathbb{E}[d_{B_{\pi}}(j)]$ and $\sigma^2 = \operatorname{Var}(d_{B_{\pi}}(j))$.

Note that the event $(\mathbb{1}_i = 1)$ only depends on the value of $\pi(i)$. There are *m* possible values in total for $\pi(i)$, and exactly the colors in which ij is an edge contribute to $\mathbb{1}$. Let $d_{H'_c}(i,j) = 1$ if $ij \in E(H'_c)$, and $d_{H'_c}(i,j) = 0$ otherwise. So

$$\Pr[\mathbb{1}_{i} = 1] = \frac{\sum_{c=1}^{m} d_{H'_{c}}(i, j)}{m}.$$

By linearity of expectations, we have

$$\mu_j = \sum_{i=1}^m \mathbb{E}[\mathbb{1}_i] = \sum_{i=1}^m \frac{\sum_{c=1}^m d_{H'_c}(i,j)}{m} = \sum_{c=1}^m \frac{\sum_{i=1}^m d_{H'_c}(i,j)}{m} = \sum_{c=1}^m \frac{d_{H'_c}(j)}{m}.$$

To compute the variance, note that for each $i \neq k$ in V_1 , we have

$$\mathbb{E}\left[\mathbb{1}_{i}\mathbb{1}_{k}\right] = \sum_{c=1}^{m} \Pr\left[\mathbb{1}_{i} = \mathbb{1}_{k} = 1 | \pi(i) = c\right] \Pr\left[\pi(i) = c\right]$$
$$= \sum_{c=1}^{m} \frac{1}{m} d_{H_{c}'}(i, j) \Pr\left[\mathbb{1}_{k} = 1 | \pi(i) = c\right]$$
$$= \sum_{c=1}^{m} \frac{1}{m} d_{H_{c}'}(i, j) \frac{\sum_{c' \neq c} d_{H_{c'}'}(k, j)}{m - 1}.$$

Thus,

$$\begin{aligned} \operatorname{Var}\left(d_{B_{\pi}}\left(j\right)\right) &= \operatorname{Var}\left(\sum_{i=1}^{m} \mathbb{1}_{i}\right) = \sum_{i=1}^{m} \operatorname{Var}\left(\mathbb{1}_{i}\right) + \sum_{i \neq k} \operatorname{Cov}\left(\mathbb{1}_{i},\mathbb{1}_{k}\right) \\ &\leq \mu_{j} + \sum_{i \neq k} \left(\mathbb{E}\left[\mathbb{1}_{i}\mathbb{1}_{k}\right] - \mathbb{E}\left[\mathbb{1}_{i}\right]\mathbb{E}\left[\mathbb{1}_{k}\right]\right) \\ &= \mu_{j} + \sum_{i \neq k} \left(\sum_{c=1}^{m} \frac{1}{m} d_{H_{c}'}(i,j) \frac{\sum_{c \neq c'} d_{H_{c'}'}(k,j)}{m-1} - \frac{\sum_{c=1}^{m} d_{H_{c}'}(i,j)}{m} \frac{\sum_{c'=1}^{m} d_{H_{c'}'}(k,j)}{m}\right) \\ &= \mu_{j} + \sum_{i \neq k} \left(\left(\frac{1}{m(m-1)} - \frac{1}{m^{2}}\right) \sum_{c=1}^{m} d_{H_{c}'}(k,j) \sum_{c'=1}^{m} d_{H_{c'}'}(i,j) - \frac{1}{m(m-1)} \sum_{c=1}^{m} d_{H_{c}'}(k,j) d_{H_{c}'}(i,j)\right) \\ &\leq \mu_{j} + \frac{1}{m^{2}(m-1)} \sum_{i,k=1}^{m} \left(\sum_{c=1}^{m} d_{H_{c}'}(k,j) \sum_{c'=1}^{m} d_{H_{c'}'}(i,j)\right) \\ &= \mu_{j} + \frac{1}{m-1} \mu_{j}^{2}. \end{aligned}$$

To complete the proof, first observe that we have $\frac{1}{m-1}\mu_j^2 \leq \frac{\alpha^2 \mu_j^2}{200}$ since *m* is sufficiently large. Also, we have $\mu_j \leq \frac{\alpha^2 \mu_j^2}{200}$ since $\mu_j \geq \frac{200}{\alpha^2}$ by assumption. Now we obtain $\sigma^2 \leq \frac{\alpha^2 \mu_j^2}{100}$ and the lemma follows.

Lemma 2.16. For every $\varepsilon > 0$ there exists $C := C(\varepsilon)$ for which the following holds for sufficiently large $m \in \mathbb{N}$ and $p = C\frac{\log m}{m}$. Let H'_1, \ldots, H'_m be bipartite graphs on the same vertex set V = [n]with the same parts $V = V_1 \cup V_2$ of the same size m, where $V_1 = [m]$. Suppose that $\delta^*(H'_c) \ge (\frac{1}{2} + \varepsilon)mp$ for every $c \in [m]$. Then for a uniformly random permutation $\pi : [m] \to [m]$, whp we have $\delta(B_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})mp$.

197 Proof. Consider B_{π} , where π is a uniformly random permutation. As $\delta^*(H'_c) \geq (\frac{1}{2} + \varepsilon)mp$ for 198 every $c \in [m]$, it is guaranteed that for all $i \in V_1$ we have that $d_{B_{\pi}}(i) \geq (\frac{1}{2} + \varepsilon)mp$. Now 199 consider some $j \in V_2$ and observe from the proof of Lemma 2.14, under the same notation, that 200 $\mu_j := \mathbb{E}[d_{B_{\pi}}(i)] \geq (\frac{1}{2} + \varepsilon)mp$.

In order to complete the proof, we want to show that the $d_{B_{\pi}}(j)$'s are "highly concentrated" using Theorem 2.4. To this end, let $h(\pi) := d_{B_{\pi}}(j)$ and note that swapping any two elements of π can change the value of h by at most 2. Moreover, note that if $h(\pi) = d_{B_{\pi}}(j) = s$, then it is enough to choose π_{proof} as the s indices reflected in $N_{B_{\pi}}(j)$. Therefore, $h(\pi)$ satisfies the conditions of Theorem 2.4 with c = 1 and r = 1. Now, let $\alpha = \frac{\varepsilon}{100}$, and observe that by Lemma 2.14 we have that the median M of $d_{B_{\pi}}(j)$ lies in the interval $(1 \pm \alpha)\mu_j$. Therefore, we have

$$\Pr\left[h \le \left(\frac{1}{2} + \frac{\varepsilon}{2}\right)mp\right] \le \Pr\left[h \le \left(1 - \frac{\varepsilon}{2}\right)\mathbb{E}[d_{B_{\pi}}(j)]\right]$$

and the latter can be upper bounded by

$$\Pr\left[h \le (1 - \frac{\varepsilon}{2})(1 + \alpha)M\right] \le \Pr\left[h \le (1 - \frac{\varepsilon}{4})M\right].$$

Now, by Theorem 2.4 we obtain that

$$\Pr\left[h \le (\frac{1}{2} + \frac{\varepsilon}{2})mp\right] \le 2\exp\left\{-\frac{(\varepsilon M/4)^2}{16M}\right\}$$

Next, using (again) the fact that $M \in (1 \pm \alpha)\mu_j$ and that $\mu_j = \Theta(mp) \ge C \log m$, we can upper bound the above right hand side by $2 \exp(-\Theta(mp)) \le n^{-2}$. Finally, in order to complete the proof, we take a union bound over all $j \in V_2$ and obtain that whp $\delta(B_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})mp$.

209 2.6. Most D_{π} 's have large minimum degree. The following lemma says that given digraphs 210 H'_1, \ldots, H'_m with large minimum semidegrees (minimum of out-degrees and in-degrees), the re-211 sulting auxiliary digraph D_{π} also has large minimum degree whp where π is a uniformly random 212 permutation.

Lemma 2.17. For every $\varepsilon > 0$ there exists $C := C(\varepsilon)$ for which the following holds for sufficiently large $n \in \mathbb{N}$ and $p = C \frac{\log n}{n}$. Let H'_1, \ldots, H'_n be graphs on the same vertex set V = [n]. Suppose that $\delta(H'_c) \ge (\frac{1}{2} + \varepsilon)np$ for every $c \in [n]$. Then for a uniformly random permutation $\pi : V \to V$, where $\delta^0(D_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})np$.

The proof of Lemma 2.17 is very similar to that of Lemma 2.16 so we leave it to the appendix.

We first give a proof of Theorem 2.10 and use it to derive Theorem 1.1.

220 Proof of Theorem 2.10. Let $\varepsilon > 0$ and $p \ge C \frac{\log n}{n}$, for a sufficiently large C. Let $m = \frac{n}{2}$. Let 221 G_1, \ldots, G_m be independent samples of B(n, p) on $V_1 \cup V_2$. Let $H_c \subseteq G_c$ be any (bipartite) subgraphs 222 with $\delta(H_c) \ge (1/2 + \varepsilon)mp$. We wish to demonstrate that whp the family of graphs H_1, \ldots, H_m has 223 a rainbow perfect matching.

Now, let π be a permutation on V_1 chosen uniformly at random. Therefore, Lemma 2.16 (with input graphs H_1, \ldots, H_m) guarantees that for almost all permutation π , $H_{\pi} = B_{\pi}(H_1, \ldots, H_m)$ (as defined in 2.7) satisfies

227 (†)
$$\delta(H_{\pi}) \ge (\frac{1}{2} + \frac{\varepsilon}{2})mp.$$

228 We focus on all π satisfying above and condition on (†).

Note that the bipartite graph $G_{\pi} = B_{\pi}(G_1, \ldots, G_m)$ is a random bipartite graph, where all pairs in $V_1 \times V_2$ are present with probability p, and independent with other pairs of vertices. Moreover, by definition, H_{π} is a subgraph of G_{π} . Thus, by (†), we have that H_{π} is a subgraph of G_{π} with $\delta(H_{\pi}) \geq (\frac{1}{2} + \frac{\varepsilon}{2})mp$. Therefore, by Lemma 2.9, whp H_{π} contains a perfect matching, which by Remark 2.8 implies that the family H_1, \ldots, H_m whp admits a rainbow perfect matching.

234 Proof of Theorem 1.1. Let $\varepsilon > 0$ and $p \ge C \frac{\log n}{n}$, for a sufficiently large C. Let $m = \frac{n}{2}$. Let 235 G_1, \ldots, G_m be independent samples of G(n, p) (on the same vertex set V = [n]). Let $H_c \subseteq G_c$ 236 be any subgraphs with $\delta(H_c) \ge (1/2 + \varepsilon)np$. We wish to show that whp the family of graphs 237 H_1, \ldots, H_m has a rainbow perfect matching.

Observe that by Lemma 2.5, whp the family of graphs G_1, \ldots, G_m satisfies

239 (‡) $(1-\varepsilon)np \le \delta(G_c) \le \Delta(G_c) \le (1+\varepsilon)np$ for all $c \in [m]$.

For the rest of the proof, we condition on (‡).

Now, let $\alpha > 0$ such that $(1 - \alpha)(1/2 + \varepsilon) \ge 1/2 + \varepsilon/2$. By Lemma 2.6 with α in place of ε , we obtain that most balanced bipartitions of $[n] = V_1 \cup V_2$ satisfy the following: for every $u \in V_i$, i = 1, 2, and for every $c \in [m]$, we have

$$d_{H_c}(u, V_{3-i}) \ge (1-\alpha) \cdot \frac{d_{H_c}(u)}{2} \ge (1-\alpha) \cdot \left(\frac{1}{2} + \varepsilon\right) \frac{np}{2} \ge \left(\frac{1}{2} + \frac{\varepsilon}{2}\right) \frac{np}{2}.$$

Now for $i \in [m]$ and any given partition $[n] = V_1 \cup V_2$, we let $G'_i := G_i[V_1, V_2]$ and $H'_i := H_i[V_1, V_2]$ be the spanning subgraphs of G_i and H_i , respectively, induced by the bipartition $V_1 \cup V_2$. Observe that, for a given partition, all G'_i are independent samples of B(n, p) on $V_1 \cup V_2$. Moreover, for most balanced bipartitions the graphs $H'_i \subseteq G'_i$ satisfy $\delta(H'_i) \ge \left(\frac{1}{2} + \frac{\varepsilon}{2}\right) \frac{np}{2}$. Therefore, by Theorem 2.10, we obtain that by taking a random partition $[n] = V_1 \cup V_2$, whp the family $\{H'_1, \ldots, H'_{n/2}\}$ admits a rainbow perfect matching. This completes the proof.

Now we give a proof of Theorem 1.2.

248 Proof of Theorem 1.2. Let $\varepsilon > 0$ and $p \ge C \frac{\log n}{n}$, for a sufficiently large C. Let G_1, \ldots, G_n be 249 independent samples of G(n, p) (on the same vertex set V = [n]). Let $H_c \subseteq G_c$ be any subgraphs 250 with $\delta(H_c) \ge (1/2 + \varepsilon)np$. We wish to show that whp the family of graphs H_1, \ldots, H_n has a 251 rainbow Hamilton cycle.

Similar to above, whp the family of graphs G_1, \ldots, G_n satisfies (‡) for m = n. For the rest of the proof, we condition on (‡).

Now, let π be a permutation on V chosen uniformly at random. Therefore, Lemma 2.17 (with input graphs H_1, \ldots, H_n) guarantees that whp $D'_{\pi} = D_{\pi}(H_1, \ldots, H_n)$ (as defined in 2.11) satisfies (*) $\delta(D'_{\pi}) \geq (\frac{1}{2} + \frac{\varepsilon}{2})np$.

257 We condition on (*).

Note that the digraph $D_{\pi} = D_{\pi}(G_1, \ldots, G_n)$ is a random digraph, where all pairs are present with probability p, and independent with other pairs of vertices. Moreover, by definition, D'_{π} is a subdigraph of D_{π} . Therefore, by Lemma 2.13, whp D'_{π} contains a Hamilton cycle, which by Remark 2.12 implies that the family H_1, \ldots, H_n whp admits a rainbow Hamilton cycle. This completes the proof.

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4. Concluding Remarks

In this note we address the Dirac-type problems for rainbow perfect matching and Hamilton cycle in a family of random graphs. Our method reduces the rainbow embedding problem to that in closely related contexts with a single host graph. That is, we use appropriate auxiliary graphs that "assemble" the family of graphs to one graph, so that the rainbow subgraph problem is reduced to finding a single copy of the desired subgraph in this auxiliary graph. For perfect matching and F-factors¹, the natural candidate for the auxiliary graph is the multi-partite graphs. For connected objects such as Hamilton cycles, we show that directed graphs are helpful auxiliary graphs.

Our method is also applicable to the dense setting and to random hypergraphs. We end by the following result for perfect matching in k-partite k-graphs. For k > d > 0 and a k-partite k-graph H, let $\delta_d^*(H)$ be the maximum integer m such that every crossing² d-set in V(H) has degree at least m. For k > d > 0, let $\delta_{k,d}$ be the smallest real number $\delta > 0$ such that for every $\varepsilon > 0$ there exists $n_0 > 0$ such that every k-partite k-graph H with n vertices in each part satisfying $\delta_d^*(H) \ge (\delta + \varepsilon)n^{k-d}$ contains a perfect matching.

¹Given graphs F and H, an F-factor in H is a spanning subgraph of H consisting of vertex-disjoint copies of F.

Theorem 4.1. Given integers k > d > 0 and $\varepsilon > 0$, there exists n_0 such that the following holds for integer $n \ge n_0$. Let H_1, \ldots, H_n be a family of k-partite k-graphs on the same k-partition with n vertices in each part. Suppose $\delta_d^*(H_i) \ge (\delta_{k,d} + \varepsilon)n^{k-d}$ for every $i \in [n]$. Then the family admits a rainbow perfect matching.

Proof Sketch. Let $V = V_1 \cup V_2 \cup \cdots \cup V_k$ be a k-partition with n vertices in each part. Given a permutation π of $V_1 = [n]$, define the auxiliary k-partite k-graph H_{π} on $V_1 \cup V_2 \cup \cdots \cup V_k$ such that a crossing k-tuple $S = \{a_1, \ldots, a_k\}$ with $a_i \in V_i$ belongs to $E(H_{\pi})$ if and only if $S \in E(H_{\pi(a_1)})$. Then for a random permutation π of V_1 , one can show that whp $\delta_d^*(H_{\pi}) \ge (\delta_{k,d} + \varepsilon/2)n^{k-d}$. Take such a permutation π and since n is sufficiently large, $\delta_d^*(H_{\pi}) \ge (\delta_{k,d} + \varepsilon/2)n^{k-d}$ implies that H_{π} contains a perfect matching M. Since M contains precisely one edge from each H_i , it is a rainbow perfect matching of the family.

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5. Appendix: Proof of Lemma 2.17

In this section, we finish the proof of Lemma 2.17 which we omitted in Section 3.

Lemma 5.1. Let $0 < \alpha < \frac{1}{2}$ and let $n \in \mathbb{N}$ be sufficiently large. Let H'_1, \ldots, H'_n be graphs on the same vertex set V = [n]. Suppose that $\delta(H'_c) \geq \frac{200}{\alpha^2}$ for all $c \in [n]$, Let π be a uniform random permutation on V and $\mu_i = \mathbb{E}\left[d_{D_n}^-(i)\right]$. Then, for every $i \in V$, we have

$$M_i := M(d_{D_{\pi}}^-(i)) \in (1 \pm \alpha) \,\mu_i.$$

Remark 5.2. The above lemma allows us to use μ_i instead of M_j in Theorem 2.4 when it is applied to $d_{D_-}^-(i)$.

- ³⁴¹ *Proof.* Consider D_{π} , where π is a uniformly random permutation on V. Let i be some vertex in V.
- Let $\mu_i = \mathbb{E}\left[d_{D_{\pi}}^-(i)\right]$ and $\sigma^2 = \operatorname{Var}(d_{D_{\pi}}^-(i))$. Moreover, for each $j \in V$, we define a random variable
- 343 $\mathbb{1}_j$, where $\mathbb{1}_j = 1$ if $\{i, j\} \in E(H'_{\pi(i)})$. Observe that $d_{D_{\pi}}(i) = \sum_{j \in V} \mathbb{1}_j$.

Applying Chebyshev's inequality, we have

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$$\Pr[|d_{D_{\pi}}^{-}(i) - \mu_i| \ge \alpha \mu_i] \le \frac{\sigma^2}{\alpha^2 \mu_i^2}$$

If we can show that $\sigma^2 \leq \frac{\alpha^2 \mu_i^2}{100}$, then the result follows. Indeed, with probability at least 99/100 we have that $d_{D_{\pi}}^-(i) \in (1 \pm \alpha) \mu_i$ and thus we conclude that the median M_i also lies in this interval. Now the remaining part is to prove the desired inequality by computing $\mu_i = \mathbb{E}[d_{D_{\pi}}^-(i)]$ and $\sigma^2 =$ $\operatorname{Var}(d_{D_{\pi}}^-(i))$.

Note that the event $(\mathbb{1}_i = 1)$ only depends on the value of $\pi(i)$. There are *n* possible values in total for $\pi(i)$, and exactly all of the colors in which ij is an edge contributes to $\mathbb{1}$. Let $d_{H'_c}(i,j) = 1$ if $ij \in E(H'_c)$, and $d_{H'_c}(i,j) = 0$ otherwise. So

$$\Pr[\mathbb{1}_j = 1] = \frac{\sum_{c=1}^n d_{H'_c}(i, j)}{n}.$$

By linearity of expectations, we have

$$\mu_i = \sum_{j=1}^{n-1} \mathbb{E}[\mathbb{1}_j] = \sum_{j=1}^{n-1} \frac{\sum_{c=1}^n d_{H'_c}(i,j)}{n} = \sum_{c=1}^n \frac{\sum_{j=1}^{n-1} d_{H'_c}(i,j)}{n} = \sum_{c=1}^n \frac{d_{H'_c}(i)}{n}$$

To compute the variance, note that for each $j \neq k$ in V, we have

$$\mathbb{E}[\mathbb{1}_{j}\mathbb{1}_{k}] = \sum_{c=1}^{n} \Pr[\mathbb{1}_{j} = \mathbb{1}_{k} = 1 | \pi(i) = c] \Pr[\pi(i) = c]$$
$$= \sum_{c=1}^{n} \frac{1}{n} d_{H_{c}'}(i, j) \Pr[\mathbb{1}_{k} = 1 | \pi(i) = c]$$
$$= \sum_{c=1}^{n} \frac{1}{n} d_{H_{c}'}(i, j) \frac{\sum_{c' \neq c} d_{H_{c'}'}(k, i)}{n - 1}.$$

Thus,

$$\begin{aligned} \operatorname{Var}(d_{D_{\pi}}^{-}(i)) &= \operatorname{Var}(\sum_{j=1}^{n} \mathbb{1}_{j}) = \sum_{j=1}^{n} \operatorname{Var}(\mathbb{1}_{j}) + \sum_{j \neq k} \operatorname{Cov}(\mathbb{1}_{j}, \mathbb{1}_{k}) \\ &\leq \mu_{i} + \sum_{j \neq k} \left(\mathbb{E}[\mathbb{1}_{j}]\mathbb{1}_{k}] - \mathbb{E}[\mathbb{1}_{j}]\mathbb{E}[\mathbb{1}_{k}] \right) \\ &= \mu_{i} + \sum_{j \neq k} \left(\sum_{c=1}^{n} \frac{1}{n} d_{H_{c}^{\prime}}(i, j) \frac{\sum_{c \neq c^{\prime}} d_{H_{c^{\prime}}^{\prime}}(k, i)}{n-1} - \frac{\sum_{c=1}^{n} d_{H_{c}^{\prime}}(i, j)}{n} \frac{\sum_{c^{\prime}=1}^{n} d_{H_{c^{\prime}}^{\prime}}(k, i)}{n} \right) \\ &= \mu_{i} + \sum_{j \neq k} \left(\left(\frac{1}{n(n-1)} - \frac{1}{n^{2}} \right) \sum_{c=1}^{n} d_{H_{c}^{\prime}}(k, j) \sum_{c^{\prime}=1}^{n} d_{H_{c^{\prime}}^{\prime}}(i, j) - \frac{1}{n(n-1)} \sum_{c=1}^{n} d_{H_{c}^{\prime}}(k, i) d_{H_{c}^{\prime}}(i, j) \right) \\ &\leq \mu_{i} + \frac{1}{n^{2}(n-1)} \sum_{j,k=1}^{n} \left(\sum_{c=1}^{n} d_{H_{c}^{\prime}}(k, i) \sum_{c^{\prime}=1}^{n} d_{H_{c^{\prime}}^{\prime}}(i, j) \right) \\ &= \mu_{i} + \frac{1}{n-1} \mu_{i}^{2}. \end{aligned}$$

To complete the proof, first observe that we have $\frac{1}{n-1}\mu_i^2 \leq \frac{\alpha^2 \mu_i^2}{200}$ since *n* is sufficiently large. Also, we have $\mu_i \leq \frac{\alpha^2 \mu_i^2}{200}$ since $\mu_i \geq \frac{200}{\alpha^2}$ by assumption. Now we obtain $\sigma^2 \leq \frac{\alpha^2 \mu_i^2}{100}$ and the lemma follows.

Proof of Lemma 2.17. Consider D_{π} , where π is a uniformly random permutation on V = [n]. As $\delta(H'_c) \geq (\frac{1}{2} + \varepsilon) np$ for every $c \in [n]$ by assumption, it is guaranteed that for all $i \in V$ we have that $\delta^+(D_{\pi}) \geq (\frac{1}{2} + \varepsilon) np$. So it suffices to prove that $\delta^-(D_{\pi}) \geq (\frac{1}{2} + \varepsilon) np$. Now consider some $i \in V$ and observe from the proof of Lemma 5.1, under the same notation, that $\mu_i := \mathbb{E} \left[d_{D_{\pi}}^-(i) \right] \geq (\frac{1}{2} + \varepsilon) np$.

In order to complete the proof, we want to show that the $d_{D_{\pi}}^{-}(i)$'s are 'highly concentrated' using Theorem 2.4. To this end, let $h(\pi) := d_{D_{\pi}}^{-}(i)$ and note that swapping any two elements of π can change the value of h by at most 2. Moreover, note that if $h(\pi) = d_{D_{\pi}}^{-}(i) = s$, then it is enough to choose π_{proof} as the s indices reflected in $N_{D_{\pi}}^{-}(i)$. Therefore, $h(\pi)$ satisfies the conditions outlined by Talagrand's type inequality with c = 1 and r = 1.

Now, let $\alpha = \frac{\varepsilon}{100}$, and observe that by Lemma 5.1 we have that the median M of $d_{D_{\pi}}^{-}(i)$ lies in the interval $(1 \pm \alpha)\mu_i$. Therefore, we have

$$\Pr\left[h \le \left(\frac{1}{2} + \frac{\varepsilon}{2}\right)np\right] \le \Pr\left[h \le \left(1 - \frac{\varepsilon}{2}\right)\mathbb{E}[d_{D_{\pi}}^{-}(i)]\right]$$

and the latter can be upper bounded by

$$\Pr\left[h \le (1 - \frac{\varepsilon}{2})(1 + \alpha)M\right] \le \Pr\left[h \le (1 - \frac{\varepsilon}{4})M\right].$$

Now, by Theorem 2.4 we obtain that

$$\Pr\left[h \le \left(\frac{1}{2} + \frac{\varepsilon}{2}\right)np\right] \le 2\exp\left\{-\frac{(\varepsilon M/4)^2}{16M}\right\}.$$

Next, using (again) the fact that $M \in (1 \pm \alpha)\mu_i$ and that $\mu_i = \Theta(np) \ge C \log n$, we can upper bound the above right band side by $2 \exp(-\Theta(np)) \le n^{-2}$. Finally, in order to complete the proof, we take a union bound over all $i \in V$ and obtain that whp $\delta^-(D_\pi) \ge (\frac{1}{2} + \frac{\varepsilon}{2})np$. 364 Department of Mathematics, University of California, Irvine. Email: asaff@uci.edu.

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