# Supplemental Materials 

# Performance Characteristics of Profiling Methods and the Impact of Inadequate Case-mix Adjustment 

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## 1 Analysis Example and R Codes

This supplemental materials document provides details on the FE flagging procedures, based on nominal p-values and empirical null adjustment, and the RE (CMS) bootstrap procedure. In addition, a sample dataset with 1,000 providers and a tutorial on the implementation along with R codes are provided for both the RE and FE methods here. The example (simulated) dataset containing 1,000 providers and 15 patient case-mix along with R functions and documentation for fitting the RE and FE models can be downloaded at www.http://faculity.sites.uci. edu/nguyen/supplement/. Note that the input dataset is sorted by provider ID's (fid).

### 1.1 Fitting RE and FE Models

The following usages fit the FE and RE/CMS models, respectively.

```
fit.FE.m1(ds, xvar.names, starting.val, numBSRuns)
```

fit.RE.CMS(ds, xvar.names, numBSRuns)

The input arguments, ds and xvar. names are the dataset name and list of patient casemix covariate names, respectively. For the FE model, starting.val is the starting value for $\gamma_{i}$ in the Newton-Raphson estimation. This can be taken to be the baseline readmission rate, for instance. The argument, numBSRuns is the number of resamples for testing the hypothesis $H_{0}: \gamma_{i}=\gamma_{M}$, which is taken to be 500 . For the RE/CMS model, numBSRuns is the number of bootstrap samples to construct the $95 \%$ confidence interval for each provider; also taken to be 500. Details are provided in the R script file (analysis_script.R) at www.http: //faculity.sites.uci.edu/nguyen/supplement/.

### 1.2 Output Objects from Fitted Models

The output object from each fitted model has two elements, the provider-level results and the patient-level case-mix coefficient estimates. For the FE model, the two elements are:
> names(fit.FE)
[1] "FE.result" "betaEstF"

The FE.result contains provider-level results $(F \times 5)$ and betaEstF contains the coefficient estimates vector $\widehat{\boldsymbol{\beta}}(r \times 1)$. The columns of FE.result are

1. sRREstsF: $S R R_{i}$ estimate, using the sum of readmissions as the numerator of SRR.
2. sRREstsF: $S R R_{i}$ estimate, using the sum of estimated probabilities of readmission as the numerator SRR.
3. pValsF: P-values from hypothesis test of $H_{0}: \gamma_{i}=\gamma_{M}, i=1,2, \ldots, F$.
4. gammaEstF: $\widehat{\gamma}_{i}$ estimate for each provider.
5. SRR.category.F.emp: Flagging indicator based on the empirical null distribution (ND: not different, B: better, W: worse).
> head(fit.FE\$FE.result)
sRREstsF sRREstsF2 pValsF gammaEstF SRR.category.F.emp
$10.93680670 .93680750 .294-1.465062$ ND
$20.81973850 .8197400 \quad 0.000-2.058169 \quad$ B
$30.84144320 .8414448 \quad 0.020-2.068698$ ND
$40.73789850 .7379005 \quad 0.000-2.416828$ B
$50.78874690 .78874850 .010-1.918674$ ND
$60.86969120 .8696927 \quad 0.044-1.639124$ ND
> head(fit.FE\$betaEstF)
[,1]
z1 0.5135744
z2 0.5241660
z3 0.5052750
z4 0.5145239
z5 0.5090906
z6 0.5069612
The output for the RE/CMS fitted model are similar, with two elements.
> names(fit.RE)
[1] "RE.result" "betaEstC"

The second element (betaEstC) contain the coefficient estimates vector $\widehat{\boldsymbol{\beta}}(r \times 1)$. The provider-level output, RE.result, consists of the five columns:

1. sRREstsc: $S R R_{i}$ estimate from the RE/CMS model.
2. sRREstsC_Lower: Lower bound of the $95 \%$ CI for $\widetilde{S R R}_{i}$ via 500 bootstrap samples.
3. sRREstsC_Upper: Upper bound of the $95 \%$ CI for $\widetilde{S R R}$.
4. BetterC: Indicator of providers flagged as "better" (1:yes; 0:no).
5. WorseC: Indicator of providers flagged as "worse" (1:yes; 0:no).black

| > | head(fit.RE\$RE.result) |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
|  | sRREstsC | sRREstsC_Lower | sRREstsC_Upper | BetterC | WorseC |
| 1 | 0.9803385 | 0.9149993 | 1.0566237 | 0 | 0 |
| 2 | 0.9034610 | 0.8284067 | 0.9758525 | 1 | 0 |
| 3 | 0.9456762 | 0.8741779 | 1.0063479 | 0 | 0 |
| 4 | 0.8969990 | 0.8147081 | 0.9803536 | 1 | 0 |
| 5 | 0.9232880 | 0.8105824 | 1.0432741 | 0 | 0 |
| 6 | 0.9386103 | 0.8620731 | 1.0212551 | 0 | 0 |

> head(fit.RE\$betaEstC)
$\left[\begin{array}{llllll}{[1]} & -0.9665475 & 0.5007279 & 0.5103138 & 0.4921882 & 0.5007380\end{array} 0.4949895\right.$

## 2 Appendix

### 2.1 RE/CMS Bootstrap Confidence Interval Estimation Procedure

We summarize here FE and RE/CMS procedures for identifying under- (over-) performing providers ("outlier" providers). The RE procedure is based on the following bootstrap CI approach, which we summarize here based on Horwitz et al. (2011) and Ash et al. (2012).

0 . Fit the generalized linear mixed effects model (1), i.e., the RE model. Denote the provider-specific estimates by $\widehat{\gamma}_{i}, i=1,2, \ldots, F$, with overall mean $\widehat{\gamma}_{0}$. Also, denote the variance and patient case-mix estimates by $\widehat{\sigma}^{2}$ and $\widehat{\boldsymbol{\beta}}$, respectively. Calculate estimated standardized readmission ratio, $S R R_{i}$ as given by (3) for the RE model.

1. Generate a bootstrap dataset by sampling $F$ providers with replacement from the original dataset. Denote the unique set of providers sampled by $\mathbb{F}^{(b)}$, where $b$ indexes bootstrap dataset.
2. Fit the RE model (1) to the bootstrap dataset and treat each resampled provider as distinct. Calculate
(a) The patient case-mix effects, $\widehat{\boldsymbol{\beta}}^{(b)}$.
(b) The mean and variance of the distribution of provider effects, $\widehat{\gamma}_{0}^{(b)}$, and $\widehat{\sigma}^{2(b)}$.
(c) The provider-specific effects and variances, $\left\{\widehat{\gamma}_{i}^{(b)}, \widehat{\operatorname{Var}}^{(b)}\left(\gamma_{i}\right)\right\}, i=1,2, \ldots, F$. (If a provider is sampled more than once, then randomly select one set of the providerspecific estimates and variances.)
3. Generate a provider random effect from the provider-specific distribution from step 2(c) for each unique provider sampled in step 1. The posterior distribution of each random effect is approximated by a normal distribution, $\widehat{\gamma}_{i}^{(b) *} \sim N\left(\widehat{\gamma}_{i}^{(b)}, \widehat{\operatorname{Var}}^{(b)}\left(\gamma_{i}\right)\right)$.
4. Calculate $S R R_{i}$ for each unique provider $i$ sampled in step 1: $S R R_{i}^{(b)}=\sum_{j=1}^{n_{i}} \widehat{p}_{i j}^{(b)} / \sum_{j=1}^{n_{i}} \widehat{p}_{M, i j}^{(b)}=$ $\sum_{j=1}^{n_{i}} g^{-1}\left(\widehat{\gamma}_{i}^{(b) *}+\widehat{\boldsymbol{\beta}}^{(b)^{\mathrm{T}}} \mathbf{Z}_{i j}\right) / \sum_{j=1}^{n_{i}} g^{-1}\left(\widehat{\gamma}_{0}^{(b)}+\widehat{\boldsymbol{\beta}}^{(b)^{\mathrm{T}}} \mathbf{Z}_{i j}\right)$, for $i \in \mathbb{F}^{(b)}$.
5. Repeat bootstrap procedure, step 1 - step 4,500 times $(b=1, \ldots, 500)$ and form the $95 \%$ confidence interval estimate of $\widetilde{S R R}_{i}$ for each provider $i=1,2, \ldots, F$.

### 2.2 FE Hypothesis Testing Procedure

The FE inference procedure is based on testing the hypothesis $H_{0}: \gamma_{i}=\gamma_{M}\left(\right.$ i.e., $\left.\widetilde{S R R}_{i}=1\right)$ for the $i$ th provider. For convenience, we summarize this procedure proposed by He et al. (2013) below. (For fitting the high-dimensional FE model using the iterative one-step NewtonRaphson algorithm, see He et al. (2013) for details.)

1. Estimate FE model parameters and fix $\boldsymbol{\beta}$ and $\gamma_{M}$ at their estimated values $\widehat{\boldsymbol{\beta}}$ and $\widehat{\gamma}_{M}$.
2. For the $i$ th provider, draw $B=500$ samples under the null hypothesis, $\left\{Y_{i j}^{(b)}: j=\right.$ $\left.1,2, \ldots, n_{i}\right\}_{b=1}^{B}$, where each sample and observations are independently drawn from a

Bernoulli distribution: $Y_{i j}^{(b)} \sim \operatorname{Ber}\left(\bar{p}_{i j}\right)$, where $\bar{p}_{i j}=\exp \left(\widehat{\gamma}_{M}+\widehat{\boldsymbol{\beta}}^{\mathrm{T}} \mathbf{Z}_{i j}\right) /\left(1+\exp \left(\widehat{\gamma}_{M}+\right.\right.$ $\left.\widehat{\boldsymbol{\beta}}^{\mathrm{T}} \mathbf{Z}_{i j}\right)$ )
3. Calculate total number of readmissions in the resampled data: $Y_{i}^{(b)}=\sum_{j=1}^{n_{i}} Y_{i j}^{(b)}$.
4. Calculate the p-value for testing $H_{0}: \gamma_{i}=\gamma_{M}$ as follows. Compute $S L_{i}^{+} \equiv B^{-1} \sum_{b=1}^{B}\left[0.5 \mathbb{I}\left(Y_{i}^{(b)}=\right.\right.$ $\left.\left.O_{i}\right)+\mathbb{I}\left(Y_{i}^{(b)}>O_{i}\right)\right]$, where $O_{i}$ is the observed number of readmissions for provider $i$ in the original data and $\mathbb{I}()$ denotes the indicator function; similarly, compute $S L_{i}^{-} \equiv$ $B^{-1} \sum_{b=1}^{B}\left[0.5 \mathbb{I}\left(Y_{i}^{(b)}=O_{i}\right)+\mathbb{I}\left(Y_{i}^{(b)}<O_{i}\right)\right]$. The p-value is $P=2 \times \min \left\{S L_{i}^{+}, S L_{i}^{-}\right\}$.
5. Repeat steps 2-4 for each provider, $i=1,2, \ldots, F$.

## 3 Supplementary Table and Figures

Table S1: Simulation settings, design parameters, and models.

| Description | Name | Parameters/summary |
| :---: | :---: | :---: |
| Provider Effect Size (P-ES) (Smaller P-ES) | P-ES 1 | $\begin{aligned} & \text { 2.5\% W: } \gamma_{i} \sim U(0.4,1.5) ; 2.5 \% \text { B: } \gamma_{i} \sim-U(0.4,1.5) \\ & 95 \% \text { ND: } \gamma_{i} \sim N\left(0,0.2^{2}\right) \end{aligned}$ |
| Provider Effect Size (P-ES) (Larger P-ES) | P-ES 2 | $\begin{aligned} & \text { 2.5\% W: } \gamma_{i} \sim U(0.6,1.5) ; 2.5 \% \mathrm{~B}: \gamma_{i} \sim-U(0.6,1.5) \\ & \text { 95\% ND: } \gamma_{i} \sim N\left(0,0.2^{2}\right) \end{aligned}$ |
| Case-mix Effect Size (CM-ES) (Smaller CM-ES) | CM-ES 1 | $\boldsymbol{\beta}_{A}: \beta_{1}=\cdots=\beta_{10}=0.5 ; \beta_{11}=\beta_{15}=1$ <br> (15 covariates) |
| Case-mix Effect Size (CM-ES) (Larger CM-ES) | CM-ES 2 | $\boldsymbol{\beta}_{B}=2 \times \boldsymbol{\beta}_{A}$ |
| CM correlation/dependence | DEP | Five cases: (1-4) $\rho=0,0.2,0.5,0.8$ for all variables; (5) general dependence/corr. structure |
| Baseline readmission rate | BRR | Low, medium, high: $14.3 \%, 27.3 \%$, and $41.7 \%$ |
| Num. of simulation studies combinations |  | $\begin{array}{r} \hline \hline 45: 3(\mathrm{P}-\mathrm{ES} 1+\mathrm{CM}-\mathrm{ES} 1, \mathrm{P}-\mathrm{ES} 1+\mathrm{CM}-\mathrm{ES} 2, \\ \mathrm{P}-\mathrm{ES} 2+\mathrm{CM}-\mathrm{ES} 1) \times 5(\mathrm{DEP}) \times 3(\mathrm{BRR}) \end{array}$ |
| Num. of providers <br> Num. of datasets |  | 1,000 per dataset 200 per combination (9,000 $=45 \times 200$ total $)$ |
| RE and FE models fitted to each dataset ( $2 \times 5=10$ models ) | $\mathcal{M}_{0}$ <br> $\mathcal{M}_{1}$ <br> $\mathcal{M}_{2}$ <br> $\mathcal{M}_{3}$ : <br> $\mathcal{M}_{f}$ | Intercept only (no case-mix adjustment) <br> Adjustment for $\left\{Z_{1}, Z_{2}\right\}$ <br> Adjustment for $\left\{Z_{1}, \ldots, Z_{10}\right\}$ <br> Adjustment for $\left\{Z_{11}, \ldots, Z_{15}\right\}$ <br> Full model with complete case-mix adjustment |
| $\overline{U(a, b)}$ denotes uniform distribution; $N\left(\mu, \sigma^{2}\right)$ denotes normal distribution. $\mathrm{W}=$ Worse, $\mathrm{B}=$ Better, and $\mathrm{ND}=$ not different relative to reference standard. |  |  |



## Case-mix Adjustment Level

Figure S1: Impact of inadequate case-mix (CM) adjustment levels (Int: Intercept only, $Z_{1}-Z_{2}$, $Z_{1}-Z_{10}, Z_{11}-Z_{15}$, Full: Full model) on RE and FE models' specificity to identify providers not different from the national reference rate for CM correlation $\rho=0.2$ across low, medium, and high baseline readmission rates (BRR). Results are presented across all providers (overall) and stratified by provider volume (large, medium, small).


Figure S2: Overall performance of the full (benchmark) RE and FE models by provider volume/size: Small, medium, and large providers defined by tertiles: (Left) Sensitivity, (Right) Specificity. Data generated were from the more general dependence (DEP) structure with unequal correlation among case-mix risk variables. For reference, given also are results for the uncorrelated and equally correlated ( $\rho=0.2$ ) case-mix scenarios.


Figure S3: Impact of inadequate case-mix (CM) adjustment levels (Int: Intercept only, $Z_{1}-Z_{2}$, $Z_{1}-Z_{10}, Z_{11}-Z_{15}$, Full: Full model) on sensitivities of RE and FE models for the more general dependence (unequal) correlation structure (DEP) among case-mix variables. Given are sensitivities to detect under-performing providers for medium baseline readmission rates. Results are presented across all providers (overall) and stratified by provider volume (large, medium, small). The case of uncorrelated case-mix is provided as a reference.


Figure S4: Overall performance of the full (benchmark) RE and FE models by provider volume (small, medium, and large) and by baseline readmission rates (BRR, low, medium and high) for simulation study based on USRDS data.


Figure S5: Impact of inadequate case-mix (CM) adjustment levels (Int: Intercept only, $Z_{1}-Z_{2}$, $Z_{1}-Z_{10}, Z_{11}-Z_{15}$, Full: Full model) on sensitivities of RE and FE models for simulation study based on USRDS data. Given are sensitivities to detect under-performing providers for medium baseline readmission rates (BRR). Results are presented across all providers (overall, row 1) and stratified by provider volume (large, medium, small - rows 2, 3, 4).


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